

BULLETIN 17B Computations

**Workshop on Determining Flood Frequencies Using Tools
from the U.S. Geological Survey**

**Presented at the 88th Annual Meeting of the
Transportation Research Board**

January 11, 2009

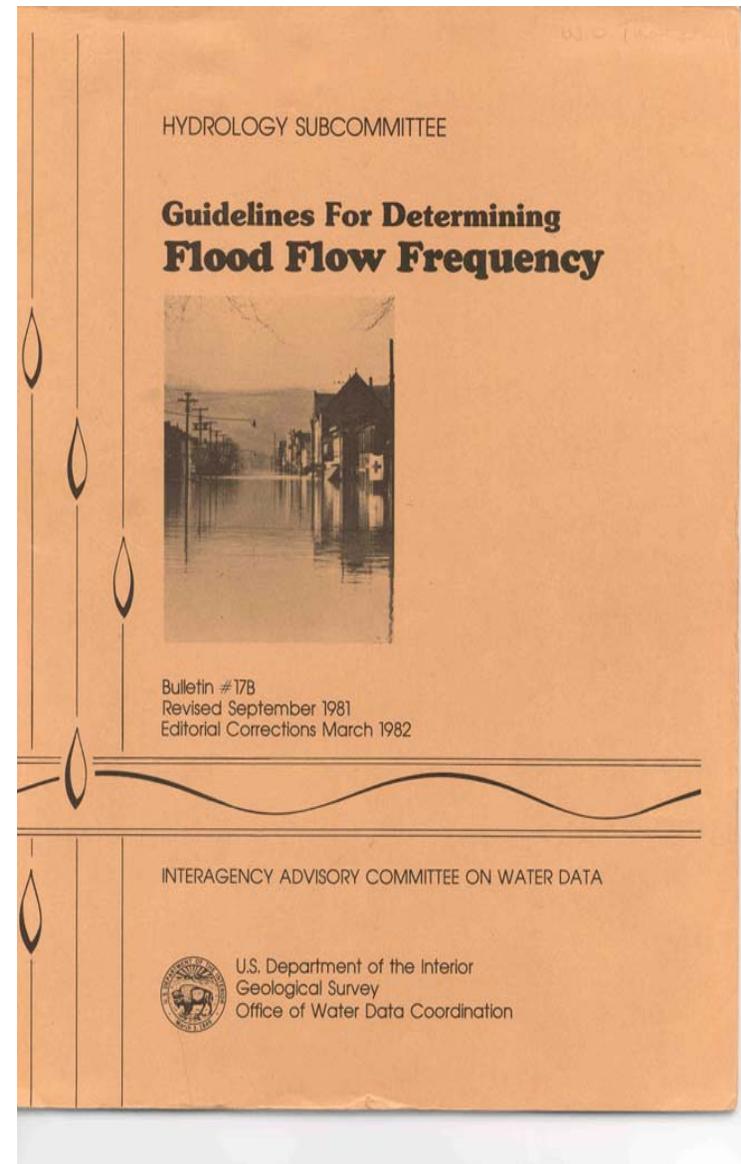
Washington, DC

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Existing Guidelines - Bulletin 17B

- **Bulletin 17B - Published in 1982, includes guidelines for:**
 - Fitting Pearson Type III distribution to logs of annual peak flows
 - Estimating generalized skew
 - Weighting generalized skew with station skew
 - Low- and high-outlier detection tests
 - Conditional probability adjustment for low outliers
 - Adjustments for historic flood data



Computing T-Year Events with the Pearson Type III Distribution

- The T-year event is computed by the method of moments as:

$$X_T = \bar{X} \pm KS$$

where X_T = T-year low flow or T-year flood event.

Computing T-Year Events with the Pearson Type III Distribution

\bar{X} = **logarithmic** mean of annual values

S = **logarithmic** standard deviation of annual values

K = The Pearson Type III frequency factor (Appendix 3 - Bulletin 17B)

Annual Maximum Peak Discharges

FLOYD RIVER AT JAMES, IOWA				
N	06600500			
H	06600500	423448	961836	1919
				882.00
3	06600500	19350628	1460	15.20
3	06600500	19360310	40502	18.101
3	06600500	19370527	3570	17.20
3	06600500	19380915	2060	16.50
3	06600500	19390312	13002	16.101
3	06600500	19400605	1390	15.40
3	06600500	19410311	1720	16.20
3	06600500	19420604	6280	18.80
3	06600500	19430617	1360	15.20
3	06600500	19440513	7440	18.80
3	06600500	19450312	5320	18.40
3	06600500	19460301	1400	15.30

<http://water.usgs.gov/nwis/sw>

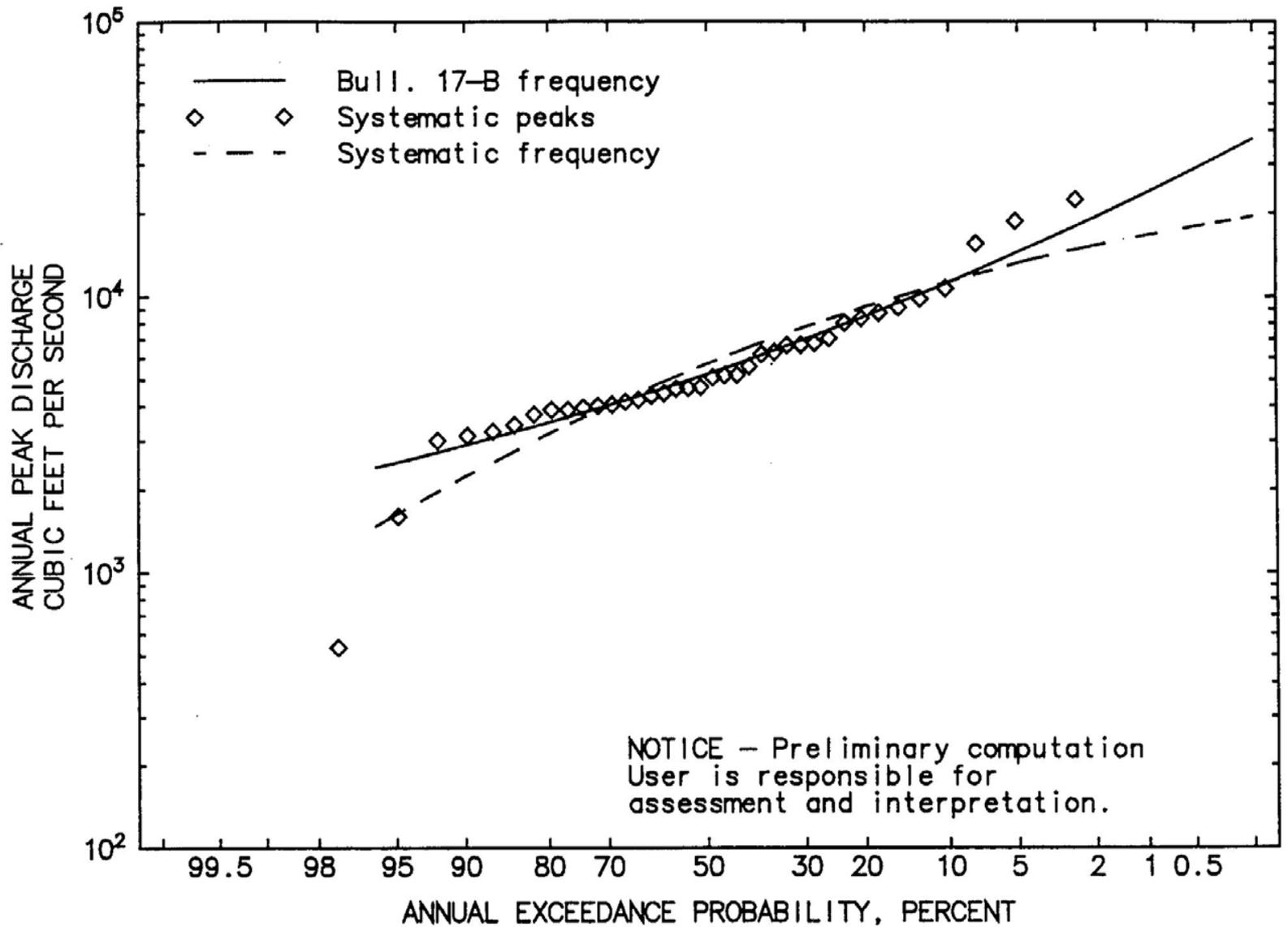
OUTLIERS AND USE OF HISTORIC FLOOD DATA

Definitions:

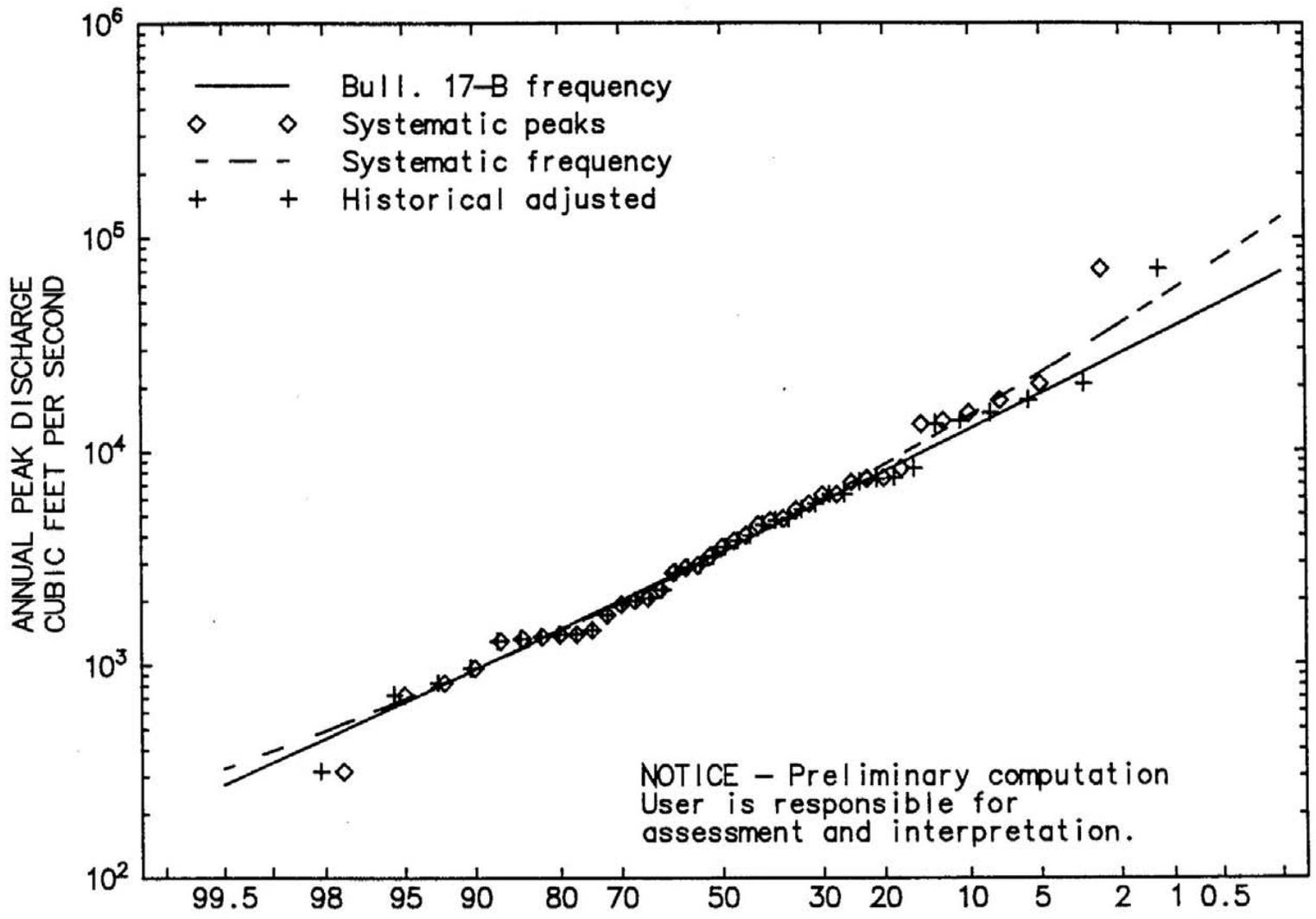
Outliers - Data points at (or near) the extremity of the frequency curve which departs from the trend of the data due to:

- measurement error problems
- statistical sampling problems

See examples of low and high outliers in the following two frequency plots.



Station - 01614000 BACK CREEK NEAR JONES SPRINGS, W.V.A.
1997 OCT 30 12:32:14



Station - 06600500 FLOYD RIVER AT JAMES, IOWA
1997 OCT 30 12:31:14

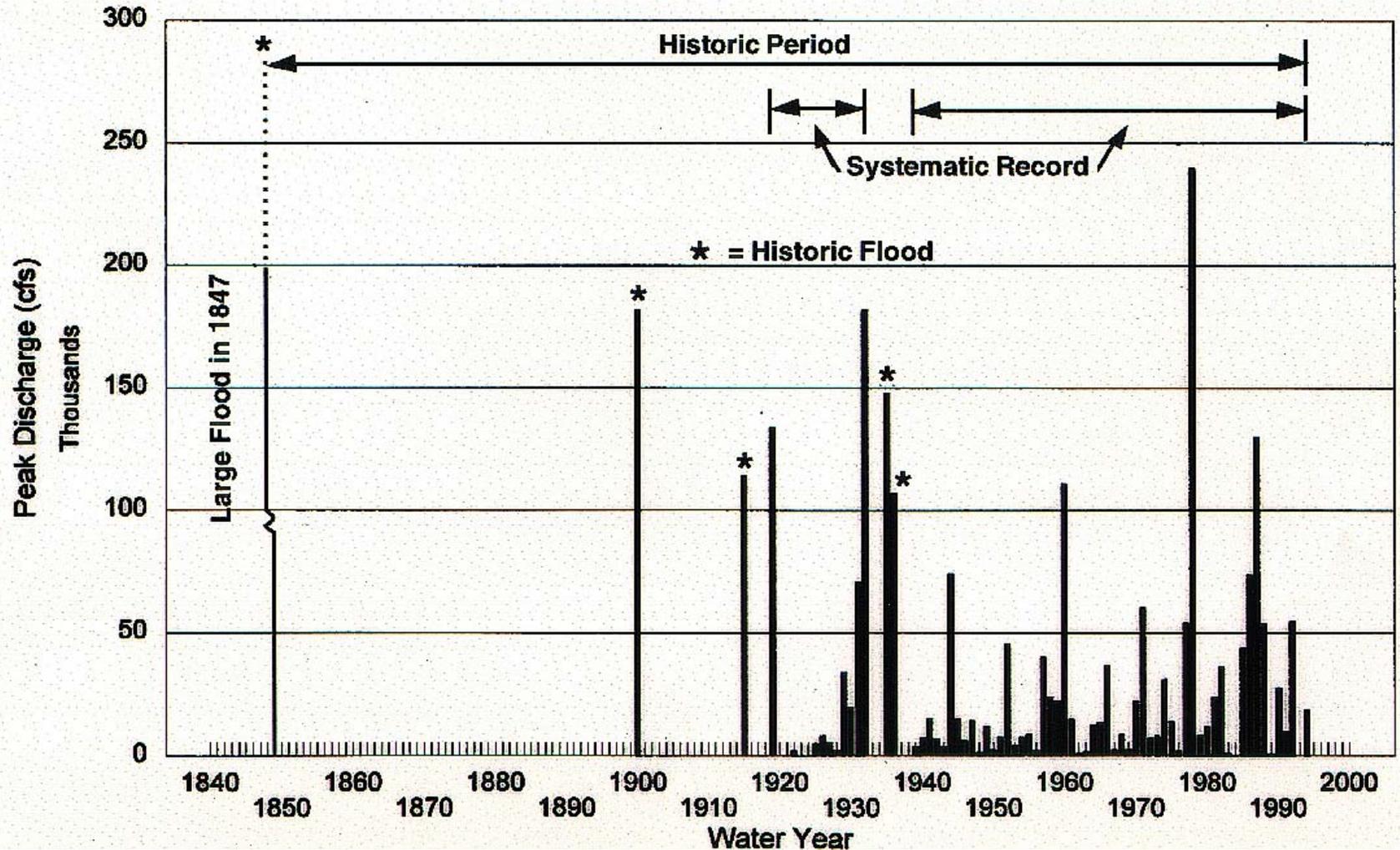
OUTLIERS AND USE OF HISTORIC DATA

Historic Data - are large events that occurred outside of the systematic record (gaged period) for which magnitudes have been **estimated**.

Historic Information - knowledge gained from residents, newspaper accounts, reports, or other information, which provide guidance on the **number of exceedances** of a large event(s) within some **specified historic period**.

Annual Maximum Peak Discharges

Guadalupe River at Comfort, TX



HISTORIC ADJUSTMENT PROCEDURE

Compute statistics of systematic and historic data.

Compute the systematic record weight,

$W = (H - Z) / (N + L)$ - defined in following figure.

H = historic period, **Z** = number of historic/high outliers, **N** = systematic years of record, **L** = number of low outliers.

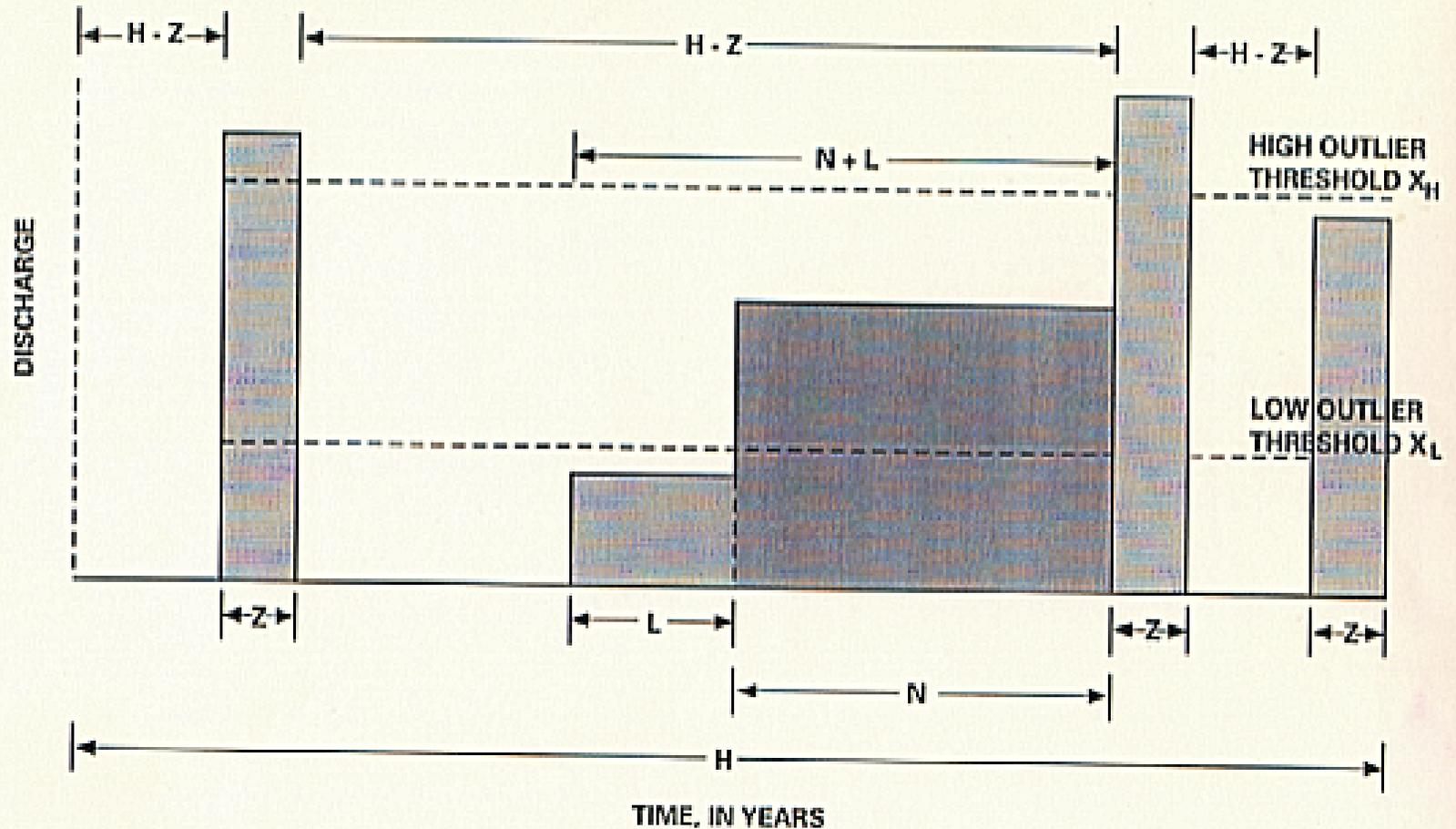
Historic peaks are given a weight of 1.

Compute the **weighted** mean, standard deviation and skew.

HISTORIC ADJUSTMENT PROCEDURE

BULLETIN 17B ADJUSTMENT FOR HISTORIC INFORMATION

$$W = \frac{H - Z}{N + L}$$



DETECTION OF OUTLIERS

High outlier (X_H) threshold value is computed by:

$$X_H = \bar{X} + K_N S$$

Where:

X_H = **high outlier** threshold (log units)

\bar{X} = logarithmic **mean** of systematic peaks excluding:

- zero flood events
- peaks below gage base
- outliers previously detected

DETECTION OF OUTLIERS

S = logarithmic **standard deviation**

K_N = K value from Appendix 4 for sample size N

Any peaks greater than threshold are considered high outliers. Historic information is needed to “adjust” the high outlier(s); if no available historic information, high outlier is left in systematic record.

*

Appendix 4

OUTLIER TEST K VALUES

10 PERCENT SIGNIFICANCE LEVEL K VALUES

The table below contains one sided 10 percent significance level K_N values for a normal distribution (38). Tests conducted to select the outlier detection procedures used in this report indicate these K_N values are applicable to log-Pearson Type III distributions over the tested range of skew values.

Sample size	K_N value						
10	2.036	45	2.727	80	2.940	115	3.064
11	2.088	46	2.736	81	2.945	116	3.067
12	2.134	47	2.744	82	2.949	117	3.070
13	2.175	48	2.753	83	2.953	118	3.073
14	2.213	49	2.760	84	2.957	119	3.075
15	2.247	50	2.768	85	2.961	120	3.078
16	2.279	51	2.775	86	2.966	121	3.081
17	2.309	52	2.783	87	2.970	122	3.083
18	2.335	53	2.790	88	2.973	123	3.086
19	2.361	54	2.798	89	2.977	124	3.089
20	2.385	55	2.804	90	2.981	125	3.092
21	2.408	56	2.811	91	2.984	126	3.095
22	2.429	57	2.818	92	2.989	127	3.097
23	2.448	58	2.824	93	2.993	128	3.100
24	2.467	59	2.831	94	2.996	129	3.102
25	2.486	60	2.837	95	3.000	130	3.104
26	2.502	61	2.842	96	3.003	131	3.107
27	2.519	62	2.849	97	3.006	132	3.109
28	2.534	63	2.854	98	3.011	133	3.112
29	2.549	64	2.860	99	3.014	134	3.114
30	2.563	65	2.866	100	3.017	135	3.116
31	2.577	66	2.871	101	3.021	136	3.119
32	2.591	67	2.877	102	3.024	137	3.122
33	2.604	68	2.883	103	3.027	138	3.124
34	2.616	69	2.888	104	3.030	139	3.126
35	2.628	70	2.893	105	3.033	140	3.129
36	2.639	71	2.897	106	3.037	141	3.131
37	2.650	72	2.903	107	3.040	142	3.133
38	2.661	73	2.908	108	3.043	143	3.135
39	2.671	74	2.912	109	3.046	144	3.138

DETECTION OF OUTLIERS

Low outlier (X_L) threshold value is computed by:

$$X_L = \bar{X} - K_N S$$

Where:

X_L = **low outlier threshold (log units)**

\bar{X} = **logarithmic mean of systematic peaks excluding:**

- **zero flood events**
- **peaks below gage base**
- **outliers previously detected**

DETECTION OF OUTLIERS

S = logarithmic standard deviation

K_N = K value from Appendix 4 for sample size N

Any peaks less than threshold are considered low outliers. When one or more low outliers are identified, they are **censored** (removed from the computations) and conditional **probability adjustment made**.

DETECTION OF OUTLIERS

If an adjustment for historic flood data has been made prior to the detection of low outliers, then use the following low outlier threshold equation:

$$X_L = \tilde{M} - K_H \tilde{S}$$

DETECTION OF OUTLIERS

Where:

X_L = low outlier threshold (log units)

\tilde{M} = historically adjusted mean
(log units)

\tilde{S} = historically adjusted standard
deviation (log units)

K_H = K value from Appendix 4 for
sample size H

When one or more low outliers are identified, they are **censored** (removed from the computations) and **conditional probability adjustment** made.

BASIS FOR OUTLIER TESTS

The K_N values in Bulletin 17B are for a one-sided 10 percent significance level test (Grubbs and Beck, 1972).

- Test for high and low outliers are made separately (one-sided)
- Hypothesis - there are no outliers
- 10 percent chance of rejecting true hypothesis

BASIS FOR OUTLIER TESTS

The test (Grubbs and Beck, 1972) was developed for detection of a **single** outlier from a **normal** distribution, but is used in Bulletin 17B for detection of multiple outliers for a Pearson Type III distribution.

Bulletin 17B Work Group evaluated various outlier tests using observed and simulated data (Thomas, 1985).

CONDITIONAL PROBABILITY ADJUSTMENT

Appropriate when annual peaks are less than the base of partial-record stations, when zero flows occur or when low outliers are identified in the record.

Basic steps in computations are:

- 1. Compute the frequency curve using the above base (or above low outlier criterion) peaks and station skew including detection of outliers and incorporation of historic information. This gives the conditional frequency curve with exceedance probabilities P_d .**

CONDITIONAL PROBABILITY ADJUSTMENT

2. Calculate the estimated probability \tilde{P} that any annual peak will exceed the truncation level

where N is the number of peaks above the truncation level and n is the total number of years of record. The equation is (Jennings and Benson, 1969):

$$\tilde{P} = \frac{N}{n}$$

CONDITIONAL PROBABILITY ADJUSTMENT

If historic information has been included, then use the following formula:

$$\tilde{P} = \frac{H - WL}{H}$$

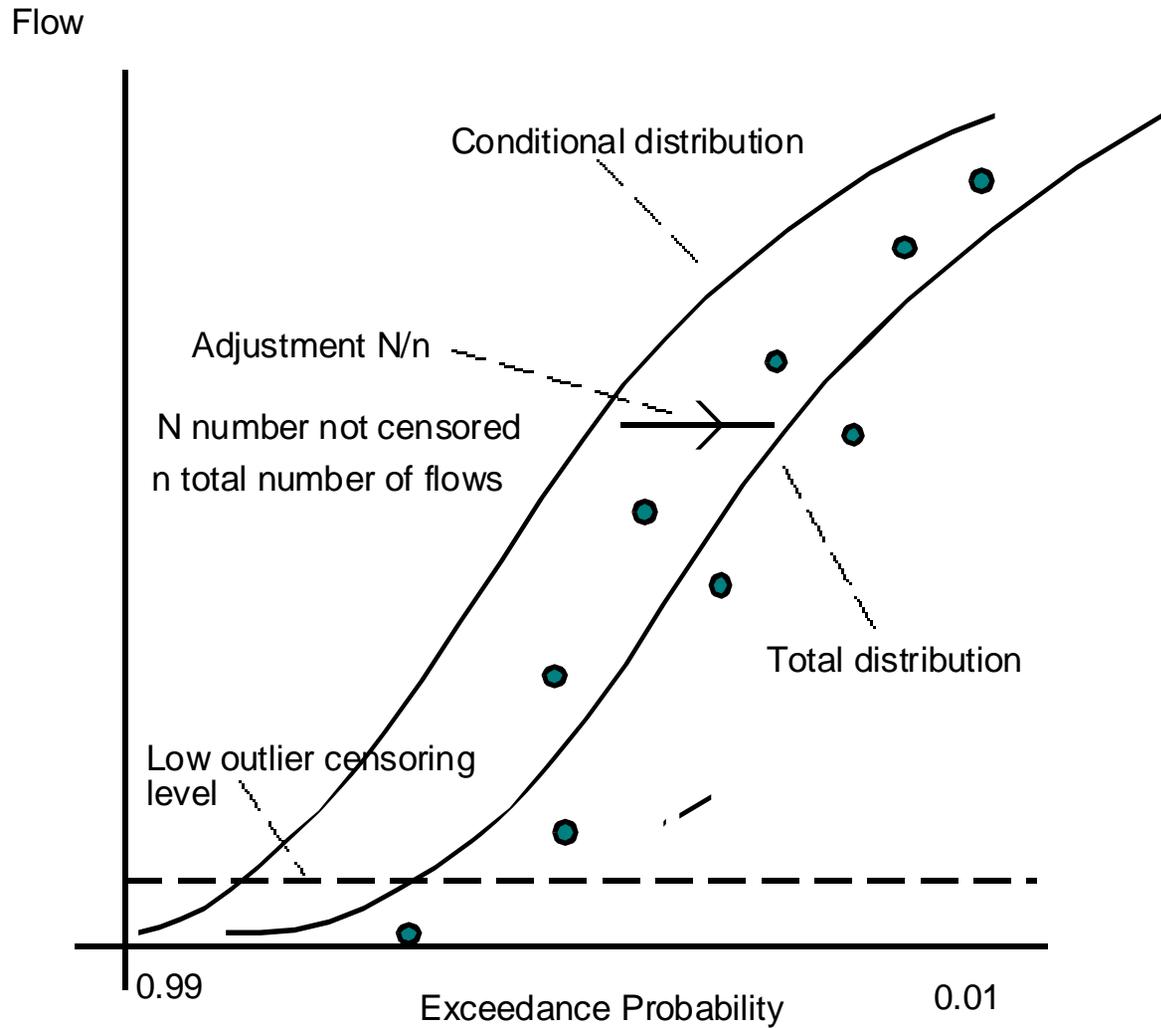
where H is the historic record length, L the number of peaks truncated, and W the systematic record weight.

CONDITIONAL PROBABILITY ADJUSTMENT

3. Adjust the frequency curve computed in 1 above for those peaks below the truncation level. This gives the unconditional frequency curve with exceedance probabilities

$$P = \tilde{P} * P_d$$

4. Interpolate either graphically or mathematically using the unconditional frequency curve in 3 above to obtain the .01, .10 and .50 exceedance probability discharges.



CONDITIONAL PROBABILITY ADJUSTMENT

5. Compute the synthetic skew (G_S) standard deviation (S_S) and mean (\bar{X}_s) using the three exceedance probabilities from 4 above and the following equations:

$$G_S = -2.30 + 3.12 \{(\log (Q_{.01}/Q_{.10}) / \log (Q_{.10}/Q_{.50}))\}$$

$$S_S = (\log (Q_{.01}/Q_{.50}) / (K_{.01} - K_{.50}))$$

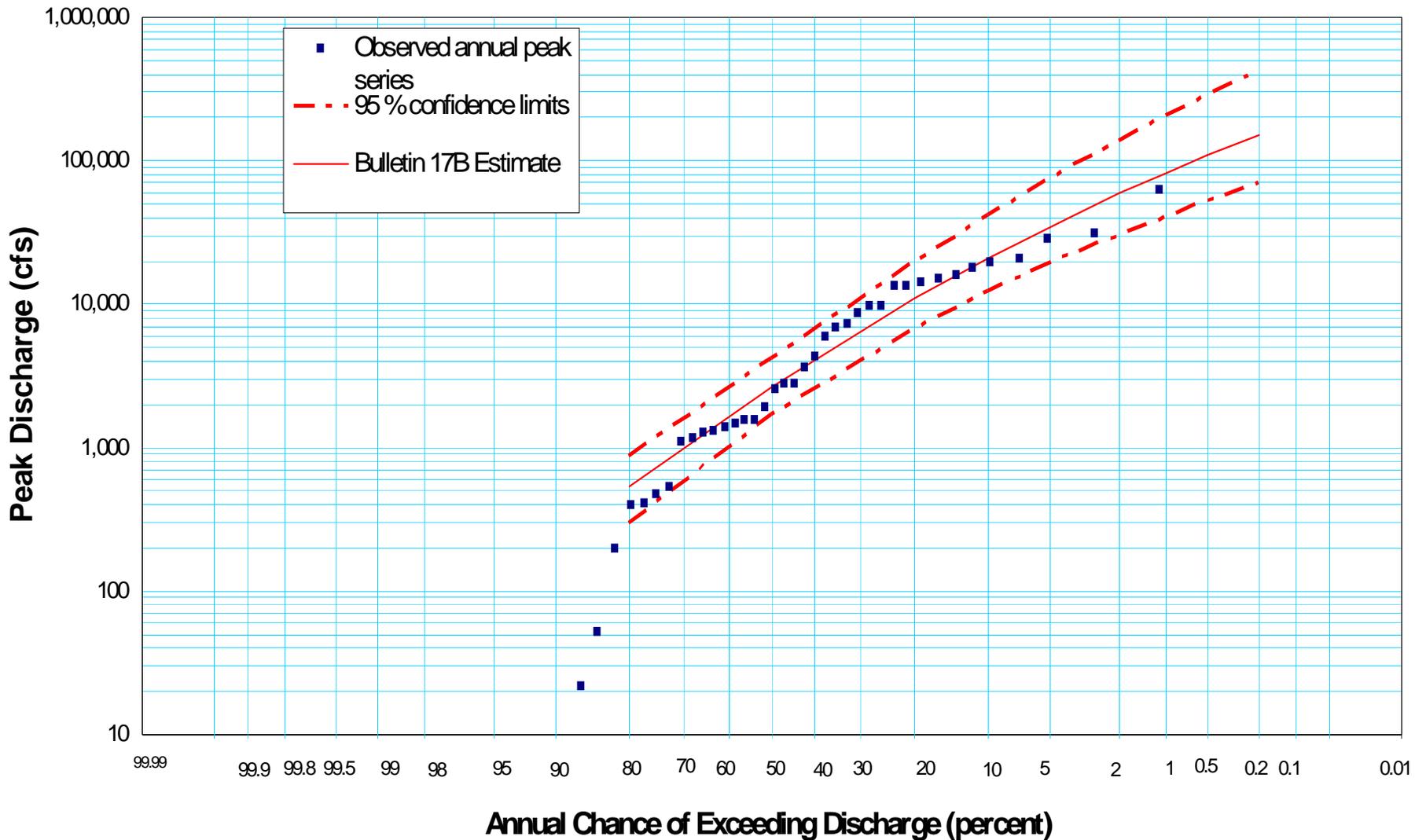
$$\bar{X}_s = \log (Q_{.50}) - K_{.50} (S_S)$$

CONDITIONAL PROBABILITY ADJUSTMENT

Where $Q_{.01}$, $Q_{.10}$ and $Q_{.50}$ are discharges with .01, .10, and .50 exceedance probabilities and $K_{.01}$, $K_{.10}$, and $K_{.50}$ are the corresponding Pearson Type III deviates.

6. The final step is to compute the weighted skew using G_S computed in 5 above and the generalized skew G .

Rocky Arroyo at Hwy BRD near Carlsbad (station 08401900)



MOTIVATION FOR GENERALIZED SKEW

Computation of coefficient of skewness is needed in Bulletin 17B method of moments approach.

There is large uncertainty in computing coefficient of skewness for sample sizes commonly available in flood-frequency analysis.

Generalized skew is used to reduce uncertainty in estimating T-year events.

PLATE I IN BULLETIN 17B

- **Map based on 2,972 stations that had 25 or more years of record through 1973 and drainage areas less than 3,000 square miles.**
- **Only 144 low outliers were found based on the Bulletin 17 (not 17B) low-outlier test that was not very sensitive.**
- **Historic flood information was not used.**
- **The national map has a mean-square error (MSE) of 0.302. Regional values of the MSE would probably be more appropriate.**

WEIGHTING THE SKEW COEFFICIENT

The weighted skew coefficient is computed as follows:

$$G_w = \frac{\text{MSE}_{\bar{G}}(G) + \text{MSE}_G(\bar{G})}{\text{MSE}_{\bar{G}} + \text{MSE}_G}$$

WEIGHTING THE SKEW COEFFICIENT

G_w = weighted skew coefficient

G = station skew

\bar{G} = generalized skew

$MSE_{\bar{G}}$ = mean square error of generalized skew

MSE_G = mean square error of station skew.

The concept of weighting the station and generalized skew in proportion to their mean square errors was based on work by Tasker (1978).

WEIGHTING THE SKEW COEFFICIENT

The mean square error of the station skew (MSE_G) can be determined from Wallis, Matalas, and Slack (1974). MSE_G can be computed as:

$$MSE_G = (\text{Bias of skew coefficient})^2 + \text{variance of skew coefficient}$$

The **bias** and **variance** of station skew coefficients for Pearson Type III random variables can be obtained from Wallis, Matalas, and Slack (1974).

WEIGHTING THE SKEW COEFFICIENT

The following equation (Wallis, Matalas, and Slack (1974)) is used for computing MSE_G as a function of **record length and skew**

$$MSE_G = 10^{[A - B [\text{LOG}_{10} (N/10)]]}$$

Where

$$A = -0.33 + 0.08 |G| \text{ if } |G| \leq 0.90$$

$$-0.52 + 0.30 |G| \text{ if } |G| > 0.90$$

$$B = 0.94 - 0.26 |G| \text{ if } |G| \leq 1.50$$

$$0.55 \text{ if } > 1.50$$

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